

1 Title

2 Increased risk of reproductive dysfunction in women prescribed long-term opioids for
3 musculoskeletal pain: a matched cohort study in the Clinical Practice Research Datalink

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- 1 The views expressed are those of the authors and not necessarily those of the NHS, the NIHR
- 2 or the Department of Health. The remaining authors have nothing to declare.
- 3 No conflicts of interest are declared
- 4 Significance: This is the first large scale cohort examining the relationship between long-term
- 5 opioid use and reproductive dysfunction using a UK national primary care database. There is
- 6 an increased risk of reproductive dysfunction associated with long-term opioid use.

1 **Abstract**

2 Background: One fifth of primary care attendees suffer chronic non-cancer pain, with
3 musculoskeletal conditions the leading cause. 12% of patients with chronic non-cancer pain
4 are prescribed strong opioids. Evidence suggests long-term opioid use is related to
5 hypogonadism in men, but the relationship in women is unclear. Our aim was to investigate
6 reproductive dysfunction in women prescribed long-term opioids for musculoskeletal pain.

7 Methods: We undertook a matched (matched 1:1; for year of birth, year of start of follow-up
8 and practice) cohort study of women aged 18-55 years old, with musculoskeletal pain and an
9 opioid prescription in the clinical practice research datalink (a primary care database)
10 between 2002 and 2013. Long-term opioid users (≥ 90 days) were compared to short-term
11 opioid users (< 90 days) for four reproductive conditions (abnormal menstruation, low libido,
12 infertility and menopause) using cox proportional hazards models.

13 Results: 44,260 women were included; the median cohort age at baseline was 43 years
14 (Interquartile Range 36-49). Long-term opioid use was associated with an increased risk of
15 altered menstruation (Hazard Ratio 1.13 95% CI 1.05 – 1.21); and with an increased risk of
16 menopause (Hazard Ratio 1.16 95% CI 1.10 – 1.23). No significant association was found for
17 libido (Hazard Ratio 1.19 95% CI 0.96 – 1.48) or infertility (Hazard Ratio 0.82 95% CI 0.64
18 – 1.06).

19 Conclusions: The risk of menopause and abnormal menstruation was increased in long-term
20 opioid users. This has implications for clinicians as reproductive dysfunction will need to be
21 considered when prescribing long-term opioids to women with musculoskeletal conditions.

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1 **Introduction**

2 Chronic Non-Cancer Pain (CNCP) can be defined as any painful condition lasting for three
3 months or more and not associated with neoplastic disease (cancer) (Chapman et al., 2010).
4 Musculoskeletal (MSK) conditions are the leading cause of CNCP (Breivik et al., 2006).

5 Over a fifth (22%) of patients attending primary care report CNCP, and women are more
6 commonly affected than men (Gureje et al., 1998). ~~In 2002, the United Kingdom (UK)~~
7 ~~primary care workload for CNCP accounted for 4.6 million General Practitioner (GP)~~
8 ~~appointments costing the National Health Service (NHS) £69 million (Belsey, 2002). In 2008~~
9 ~~the annual cost of CNCP in the United States was between \$560 and \$635 billion, an~~
10 ~~economic burden greater than for heart disease, diabetes or cancer (Gaskin and Richard,~~
11 ~~2012).~~

12 Opioid prescribing has been increasing over the past 20 years and guidelines recommend
13 their use as a second line treatment in MSK pain (Cheung et al., 2014; National Institute for
14 Health and Care Excellence, 2014; World Health Organisation, 1990). A cross-sectional
15 study in Europe found 12% of those with CNCP in the UK used strong opioids and 50% used
16 weak opioids (Breivik et al., 2006). A ~~large American cohort found an annual increase in~~
17 ~~long term opioid prescribing between 1997-2005 of 6%, and a more~~ recent British cohort
18 found that the number of people starting a long-term opioid increased by 38% between 2002
19 and 2009 (Bedson et al., 2016). Evidence for opioids' effectiveness in CNCP is weak (Els et
20 al., 2017a; Eriksen et al., 2006; Higgins and Green, 2011; Kissin, 2013; Noble et al., 2010),
21 and there is a growing body of evidence that they can have significant side effects, including
22 increased morbidity, mortality and overdose risk (Dunn et al., 2010; Els et al., 2017b;
23 Saunders et al., 2009).

24 Recently, concerns surrounding opioids and their effects on the endocrine system have arisen.
25 ~~Guidelines from the~~ The British Pain Society highlight the likelihood of endocrine side effects

1 from opioids (including reproductive dysfunction), but they conclude that there still remains
2 insufficient data to be able to quantify the risk, ~~associated with long-term opioids and a recent~~
3 Cochrane review found no current evidence for hypogonadism or sexual dysfunction (Els et
4 al., 2017b; The British Pain Society, 2010). Current literature suggests that long-term opioid
5 therapy has ~~an strong~~ impact on the male reproductive system, for which the primary
6 mechanism is thought to be through suppression of the hypothalamic-pituitary-gonadal
7 (HPG) axis which leads to low testosterone levels and reproductive and sexual dysfunction
8 (Abs et al., 2000; Aloisi et al., 2009; Benyamin et al., 2008; Daniell, 2002; Katz and Mazer,
9 2009; Smith and Elliott, 2012). Potentially, opioids may have similar biochemical and
10 clinical effects in women, with the addition of altered menstrual cycles and possibly
11 galactorrhoea (Ballantyne and Mao, 2003; Brennan, 2013; Brown and Zueldorff, 2007;
12 Colameco and Coren, 2009; Katz and Mazer, 2009; Schmittner et al., 2005; The British Pain
13 Society, 2010). A systematic review examining reproductive dysfunction and prescribed
14 long-term opioids in women found 12 relevant papers, and although they were small studies,
15 the majority suggested a link between long-term opioid use and reproductive dysfunction,
16 both biochemically (decreased sex hormones) and clinically (decreased libido and altered
17 menstrual cycle) (Wersocki et al., 2017).

18 The aim of this study was to assess if long-term opioid use for MSK pain, compared to short-
19 term use is associated with reproductive dysfunction in women.

20 **Methods**

21 *Study setting and population*

22 This study used data from the Clinical Practice Research Datalink (CPRD). CPRD is a high-
23 quality, anonymised, large UK primary care database containing information on over 11
24 million patients from over 600 primary care practices; there are currently between 4 and 5
25 million active patients (Herrett et al., 2015; Williams et al., 2012). UK primary care databases

1 provide almost complete population coverage, as over 98% of the population are registered
2 with GPs, who are the gatekeepers to secondary care (Garcia-Rodriguez and Perez
3 Gutthann, 1997; Herrett et al., 2015; Williams et al., 2012). CPRD has been compared with
4 the UK census and was shown to be broadly representative of the UK population in terms of
5 age, ethnicity and sex (Herrett et al., 2010). CPRD includes data on patients' clinical
6 conditions and prescribed medications. For the purpose of analysis, the practices included (n
7 = 350, all from English regions) were those linked to the Office for National Statistics (ONS)
8 neighbourhood deprivation data (the Index of Multiple Deprivation) and Hospital Episode
9 Statistics (HES) (Department for Communities and Local Government, 2011). Practices with
10 linked data have been shown to be similar to practices without linkage in respect of
11 demographic data, years of follow-up and prescribing of medication (Gallagher et al., 2011).
12 All doctors in practices contributing to CPRD are trained in their recording of consultations
13 and diagnostic codes and CPRD has been validated in previous studies (Herrett et al., 2010,
14 2015).

15 *Study participants*

16 This study represents secondary data analysis of an established cohort (Bedson et al., 2016).
17 We identified women aged 18 to 55 years old from within the cohort, starting a long-term
18 opioid (defined below) and with a coded non-inflammatory potentially painful MSK
19 condition from 2002-2012. The following factors were used as exclusion criteria: a cancer
20 diagnosis at any time prior to the first day of opioid use or within the following six month and
21 less than one year of records within CPRD prior to the first day of opioid prescription. MSK
22 conditions were identified from the database using a previously defined list of Read codes (a
23 clinical coded dictionary used for recording consultations in UK primary care IT systems; list
24 available online at www.keele.ac.uk/mrr) (Health and Social Care Information Centre, 2015;
25 Jordan et al., 2010). Opioids were defined as analgesics used to relieve moderate or severe

1 pain and were identified from sections 4.7.1. and 4.7.2. of the British National Formulary
2 (BNF) (BNF, 2014). This definition includes weak opioids such as codeine, and potent
3 opioids such as morphine, both in long-acting (duration of action 12 hours or longer) and
4 short-acting (duration of action 4-6 hours) forms (Fallon et al., 2006). Long-term opioid use
5 was defined as the issue of least three opioid prescriptions within 90 days from and including
6 the first date of a new prescription for opioids. A new episode of opioid use starts on the date
7 of an opioid prescription if there had been no opioids prescribed within the previous six
8 months. This definition is in line with classification of long-term opioid use in previous
9 studies and the definition was developed by Von Korff et al (2008) following exploratory
10 analyses (Bedson et al., 2016; Dunn et al., 2010; Von Korff et al., 2008). Any opioid
11 prescription not fitting this definition at any point during 2002-2012 was defined as short-
12 term opioid use (i.e. a maximum of two prescriptions within a 90 day period). The end of a
13 period of opioid use was defined by a gap of more than six months from last use of opioids,
14 which was taken to be 28 days after the issue of the last prescription (prescribing guidelines
15 for controlled drugs from the NHS Business Services Authority state that no more than a 28
16 day supply should be given except in exceptional circumstances)(N H S Business Services
17 Authority, 2014).

18 MSK conditions were identified from between 14 days before and 90 days after the start of a
19 new episode of opioid use in order to create a temporal association between a new opioid
20 prescription and an MSK condition. MSK conditions were chosen as they represent a large
21 proportion of CNCP patients, with arthritis being identified as the cause for CNCP is 40% of
22 UK patients. MSK conditions are a homogenous group and limit confounding by indication
23 (Breivik et al., 2006). Inclusion and exclusion criteria are detailed in Table 1.

24 *Study design*

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1 We conducted a matched cohort study consisting of women starting a long-term opioid
2 (exposed population) at the time of a MSK condition, matched to women starting a short-
3 term opioid (non-exposed population) at the time of a MSK condition. We matched for year
4 of birth (up to ± 5 years), general practice and first year of opioid use (up to ± 2 years).

5 *Follow-up*

6 Follow-up started from 90 days after the initial opioid prescription (in both short-term and
7 long-term opioid users) until the study outcome occurred, or for a maximum of 5 years.
8 Censoring occurred if the patient no longer contributed data (death, practice leaving CPRD or
9 patient leaving a CPRD practice) or if the study end date was reached (10th January 2012)
10 prior to the end of the 5 year follow-up period.

11 *Outcome Measures*

12 The primary outcomes of interest were defined following a systematic review by the authors,
13 these were altered menstruation, amenorrhoea (absent menstruation) or oligomenorrhoea (less
14 frequent menstruation, either less than 9 menstrual periods a year or at least 35 days in
15 between menstrual periods), decreased libido, menopause and infertility (Impey and Child,
16 2012; Norwitz and Schorge, 2013; The Practice Committee of the American Society for
17 Reproductive Medicine, 2008; Wersocki et al., 2017). These outcomes were identified
18 through searching the database for relevant Read codes from 90 days after commencing
19 opioids until end of follow-up. The read code list for each outcome was developed by EW
20 and then reviewed by JB who is experienced in developing read code lists and a consensus
21 decision was reached; both are academic GPs in the UK (lists available online at
22 www.keele.ac.uk/mrr). Outcomes were analysed separately so each participant could have
23 more than one outcome.

24 *Covariate data*

1 Data on comorbidities was collected over a 15-month time period, from 12 months before
2 initial opioid prescription until 3 months after (when follow-up began). Three types of
3 comorbidity data were identified. The outcomes of interest were considered comorbidities if
4 they occurred prior to the start of follow-up, and they were identified in the same way as
5 described for outcome measures. The following covariates were also identified: thyroid
6 conditions, low BMI <18 (as a coded condition), adrenal conditions, obesity (as a coded
7 condition), and BMI (categorised as <25 kg/m², ≥25 kg/m² (overweight) or missing) was
8 recorded at the date closest to the start of follow-up, structural gynaecology condition and
9 illegal opioid misuse. These conditions were chosen following a literature search for common
10 conditions that could be associated with the outcomes of interest. Finally the total number of
11 prescriptions were mapped to BNF sections and used as a surrogate measure for the number
12 of comorbid conditions (Perkins et al., 2004).

13 Alcohol and smoking were categorised as never, ever or missing and was identified prior to
14 the start of follow-up. Age was included within analysis as a continuous variable.

15 Non-steroidal anti-inflammatory drugs (NSAIDs) have been associated with anovulation,
16 therefore we specifically identified co-prescribing with NSAIDs for 4 months prior to the
17 start of follow-up (Salman et al., 2015).

18 *Statistical Analysis*

19 The baseline demographics of the two groups were first described and compared using
20 appropriate simple statistical tests (Wilcoxon-Mann-Whitney or Chi squared test).

21 The rates of adverse events per 10,000 person-years at risk in short-term and long-term
22 opioid user groups were determined. Cox proportional hazards models were used to produce
23 both unadjusted and adjusted hazard ratios (HRs), estimating the effects of association
24 between opioid use and reproductive dysfunction. The validity of the proportional hazards

1 assumption was tested using Schoenfeld residuals, where the proportional hazards
2 assumption was violated those covariates were included as time varying covariates (see Table
3 [34](#)) (Bellera et al., 2010). HRs were adjusted for the covariates described above. Sensitivity
4 analyses were undertaken using first a complete case approach, secondly removing those with
5 pre-existing outcome conditions and finally removing all those women with reported
6 menopausal symptoms (a proxy for analysis of those not receiving Hormone Replacement
7 Therapy (HRT)). Data was analysed using STATA 14.0 (StataCorp, 2017).

8 *Ethics committee approval*

9 Approved by the CPRD Independent Scientific Advisory committee (reference number
10 ISAC, Protocol No. 13_135).

11 **Results**

12 The study cohort included 44,260 women (22,130 long-term opioid users and 22,130 short-
13 term users). Table [12](#) shows the baseline characteristics of the two groups. The median age in
14 both groups was 43. Long-term opioid users were more likely to use NSAIDs, have a higher
15 BMI, have more comorbidities and to be smokers when compared with short-term opioid
16 users (see Table [12](#)). Median follow-up was 39 months, the total number of years under
17 follow-up for each outcome is shown in table [23](#).

18 The incidence of abnormal menstruation was higher for long-term opioid users (209.5 per
19 10,000 person-years (95% CI 199.0 – 220.7)) than for short-term opioid users (186.2 per
20 10,000 person-years (95% CI 176.4 - 196.7)). The adjusted HR showed an increased risk of
21 abnormal menstruation 1.13 (95% CI 1.05 – 1.21) for long-term opioid users compared to
22 short-term opioid users over the five year follow-up.

23 The incidence of menopausal symptoms was higher in long-term opioid users (383.7 per
24 10,000 person years (95% CI 369.1 – 398.8)) than in short-term opioid users (330.9 (95% CI

1 317.5 – 344.8)). The adjusted HR was 1.16 (95% CI 1.10 – 1.23) for long-term users
2 compared with short-term users.

3 Low libido had an incidence in long-term opioid users of 27.7 per 10,000 person years (95%
4 CI 24.0 – 31.8) and in short-term users of 22.6 (95% CI 19.4 – 26.4). The adjusted hazard
5 ratio did not find a statistically significant difference in risk of low libido between long-term
6 and short-term opioid users (adjusted HR 1.19, 95% CI 0.96-1.48).

7 Infertility had an incidence in short-term opioid users of 19.0 per 10,000 person years (95%
8 CI 16.0 – 22.5) and a lower incidence in long-term opioid users of 16.4 (95% CI 13.7 – 19.7).
9 However, there was no statistically significant difference in risk of infertility seen with the
10 adjusted Cox regression between short-term and long-term opioid users (adjusted HR 0.82,
11 95% CI 0.64 – 1.06).

12 Sensitivity analysis using only complete cases, [removing those with menopausal symptoms](#)
13 and also for those without pre-existing reproductive dysfunction gave similar results (see
14 Table [45](#)).

15 **Discussion**

16 To our knowledge, this study of over 40,000 opioid users in primary care is the first to
17 demonstrate an increased risk of abnormal menstruation and of menopausal symptoms in
18 women aged 18-55 years prescribed long-term opioids compared with short-term opioids.
19 The relationship between opioid use and low libido and infertility was not so clear. These
20 results are consistent with the limited current literature, where a recent systematic review
21 found a possible relationship between opioid use and female reproductive dysfunction
22 (Wersocki et al., 2017).

23 Women aged 18-55 years old with painful MSK conditions and prescribed opioids were
24 identified for this study. We used MSK pain to define the cohort rather than all cause CNCP

1 as MSK conditions represent a large proportion of CNCP cases and are a more homogenous
2 group of conditions, in order to address some elements of indication bias (Breivik et al.,
3 2006). However, residual confounding cannot be ruled out as one explanation for the
4 findings, particularly where severity and longevity of the MSK condition is not taken into
5 account. We also cannot be completely confident that the opioid was prescribed for an MSK
6 condition, since if the indication was a different type of pain this may introduce additional
7 confounding. As our cohort included only participants with MSK pain this may mean the
8 results are not generalizable to other types of CNCP. It is also important to note that the MSK
9 conditions include all joints and all MSK conditions that would not be considered
10 inflammatory. Accordingly, this still represents a broad group of conditions.

11 The women in the long-term opioid use group also had a higher number of comorbidities than
12 the short-term opioid use group which may indicate that they may experience poorer health in
13 general. Additionally the increased stress levels that may be associated with increased
14 comorbidity can be associated with some of the outcomes of interest; however, this was
15 adjusted for during Cox regression (The Practice Committee of the American Society for
16 Reproductive Medicine, 2008). Long-term opioid users were also more likely to be smokers
17 and there is some evidence that smoking can be associated with premature ovarian failure
18 (POF), however we were able to include smoking within the analysis in order to account for
19 this (Luborsky et al., 2003). There was a significant difference seen for alcohol use as well
20 with long-term opioid users appearing to drink less alcohol. There is evidence that alcohol
21 use can be associated with low libido and this was also adjusted for within the analysis
22 (Arunakumari and Walker, 2009). The groups were also significantly different in statistical
23 terms with respect to ethnicity where there was also a large amount of missing data, due to
24 the missing data. As a consequence we were unable to adjust for ethnicity and this is a

1 potential limitation as there is some evidence that distinct ethnic groups enter the menopause
2 at different ages and experience variation in rates of POF (Luborsky et al., 2003).

3 Everyone in the cohort was considered suitable to receive prescribed opioids and this
4 decreases potential systematic differences between the comparison groups. However, the
5 groups may still be different in terms of the severity or nature of the underlying condition.
6 We have also not assessed the impact of daily morphine equivalent opioid dose and it may be
7 that long-term users were receiving higher daily doses, although there is evidence in the UK
8 that the majority of long-term opioid users are receiving less potent opioids such as codeine

9 15mg, this would be an important area for future research (Bedson et al., 2016). The
10 identification of women as either long-term or short-term opioid users was based on issued
11 prescriptions, although we cannot definitely establish how or if the patient used the
12 medication. It is also important to note that opioid exposure was only identified at baseline
13 and participants did not move in and out of opioid exposure groups, whereas in real life
14 opioid exposure is time-varying so this may have an impact on the generalisability of the
15 results. The study is unable to report on the risk associated with opioid use compared to no
16 opioid use for the adverse effects studied (as the comparison was between opioid duration),
17 this is a limitation of the study. However it was felt that comparing duration of opioid use
18 reduced the likelihood of confounding by indication and this was an important consideration
19 when designing the study. It also removed the possibility of contraindications to opioid
20 prescribing affecting the prescription of opioids in either group. Further attempts were made

21 to control for confounding through statistical adjustment for important factors including: age,
22 BMI, smoking status, NSAID use and medical conditions associated with the outcomes of
23 interest. The study did not identify women receiving HRT or hormonal contraception and
24 these may have affected results, but there is no reason to suspect usage is different between
25 the short-term and long-term opioid users. We undertook a sensitivity analysis where women

1 with a diagnosis of menopause were removed and there was no difference seen in the results
2 for the other outcomes (see Table 4). We were unable to include hormonal contraception
3 within the analysis as CPRD does not capture all the relevant information, in the period from
4 April 2012 to March 2013 1.2 million women attended family planning clinics and of these
5 47% received oral contraception and this information would not be routinely recorded within
6 CPRD (Health and Social Care Information Centre, 2013). Given the different routes through
7 which women can access oral contraception we were unable to confidently include this
8 within the analysis.

9 The use of CPRD was a strength, as it allowed us access to a large sample of opioid users
10 which provided enough statistical power to undertake the appropriate analysis. A particular
11 strength of this study when compared to the current body of evidence which is mainly set in
12 secondary care is its generalisability to primary care, as CPRD is broadly representative of
13 the UK population (Herrett et al., 2010). Only 23% of pain patients are seen by a pain
14 specialist, so these secondary care studies may represent a subset of patients with more severe
15 pain conditions (Breivik et al., 2006). There will be little missing data, as in the UK
16 prescriptions are recorded automatically in the electronic records on issue and CPRD
17 practices must code a condition when prescribing; this will also have helped reduce missing
18 data when identifying outcomes of interest (Lawson et al., 1998). In the UK the only opioids
19 available without a prescription are low dose codeine and dihydrocodeine and then they are
20 only recommended for a maximum of three days for acute painful injuries and come in a
21 maximum pack size of 32 tablets, so any long-term opioid use should be via prescription
22 (Medicines and Healthcare products Regulatory Agency, 2009). CPRD has previously been
23 shown to be valid for identifying patients with MSK conditions (Jordan et al., 2006). A
24 review of CPRD studies found a median proportion of confirmed diagnoses of 89% over 357
25 validation studies examining 183 diagnoses (Herrett et al., 2010).

1 The number of women with low libido within the cohort was lower than would be expected,
2 with only 0.8% of the cohort having a coded diagnosis compared to population estimates of
3 between 25 and 50% (Dunn et al., 1998; Laumann et al., 2005). This is likely to have affected
4 the possibility of finding a statistically significant relationship. This also raises the question
5 as to why there were such low numbers. Sexual dysfunction is more subjective condition in
6 women when compared to men and they may worry about their privacy in relation to
7 presenting clinically (Montgomery, 2008). Treatment options for sexual dysfunction in
8 women is poor. Consequently women with sexual dysfunction infrequently present in
9 primary care creating a clinical iceberg (Last, 1963). Investigation of low libido might be
10 better undertaken in other ways, for instance a cross-sectional study.

11 In conclusion, this large cohort study has uniquely found that women aged 18-55 years
12 receiving long-term opioids for painful MSK conditions were at increased risk of menopause
13 and abnormal menstruation when compared with women receiving short-term opioids. This is
14 an important finding against the background of increasing long-term opioid use (Bedson et
15 al., 2016). This should be considered when clinicians are prescribing opioids and discussed
16 with patients. If a patient is receiving opioids, ~~it should also~~ this should be considered ~~as a~~
17 ~~possible causative factor~~ when investigating women presenting with reproductive
18 dysfunction. It is recommended that clinicians undertake regular reviews of patients
19 prescribed long-term opioids and to assess the effectiveness of the medication but also to
20 raise the possibility of adverse effects including reproductive adverse effects as patients may
21 not volunteer these symptoms. The study was unable to find a definite link for low libido,
22 and further research is required, potentially using direct patient data collection. Further
23 research may also be needed to identify whether there is a daily dosage level where risk
24 becomes more significant.

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9 or the Department of Health. The remaining authors have nothing to declare.

10 **Author Contributions**

11 JB, YC, ER and KD were involved in the development of the research question and methods.

12 ER and YC had full access to all the data in the study and take responsibility for the integrity
13 of the data and the accuracy of the data analysis.

14 All authors discussed and interpreted the results, commented on the manuscript and approved
15 the final draft.

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16 | Table 1. Characteristics of study participants on short- and long-term opioids (n=44260).

17 | Table 1 footnote: Data is either presented as n,% or median (interquartile range). P value was
18 obtained from Wilcoxon-Mann-Whitney or Chi-squared tests where appropriate.

19

20 | Table 2. Hazard Ratios for the relationship between length of opioid exposure and each
21 reproductive dysfunction outcome.

22 | Table 2 footnote: ^a adjusted for the outcome of interest (if pre-existing), thyroid conditions,
23 structural gynaecological conditions, illegal opioid use, NSAID use, BMI (<25kg/m²,
24 ≥25kg/m² or missing), smoking status, alcohol use and age

25

26 | Table 3 Proportional-hazards assumption test and time-varying covariates entered into each
27 Cox regression model.

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3 | Table 4 Sensitivity Analysis for complete cases only, excluding women with menopausal
4 | symptoms (both pre-existing and new) and if outcome is present prior to start of follow-up

