### Introduction

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- 2 Osteoarthritis (OA) is a common reason for consultation in primary care with 4% of adults
- aged 45 years and over consulting each year for OA (Jordan et al. 2013).
- 4 Although traditionally viewed as a degenerative disease of joints, OA can be considered to
- 5 have different phenotypes of disease with distinct clinical characteristics or causal factors
- 6 (Bijlsma, Berenbaum & Lafeber 2011). Obesity is a risk factor for OA (Cooper et al. 1998) but
- 7 this may be more than a purely mechanical effect (Sellam, Berenbaum 2013). Diabetes has
- 8 been associated with different musculoskeletal conditions and has also been identified as a
- 9 risk factor for OA, independently of body mass index (BMI) (Louati et al. 2015). Conversely
- an increased risk of type 2 diabetes in people with OA has also been identified (Rahman et
- al. 2014). This OA phenotype has been described as an expression of the metabolic
- syndrome (Velasquez, Katz 2010). Insulin resistance is thought to be associated with the
- development of OA (Hamada et al. 2015).
- 14 Statin treatment for primary or secondary prevention of vascular disease has been found to
- be associated with a reduction in some manifestations of clinical OA (Kadam, Blagojevic &
- 16 Belcher 2013). Although a causal mechanism for this association has not been established, it
- is plausible that the relationship is due to OA forming a part of the metabolic syndrome.
- 18 Metformin as a treatment for type 2 diabetes has previously been investigated to determine
- whether it is associated with a reduction in the risk of cardiovascular events and all-cause
- 20 mortality, but with conflicting findings (Boussageon et al. 2012).
- 21 There is limited knowledge about the effects of metformin on risk of OA. It has been
- 22 hypothesised that metformin is associated with bone health through promotion of

- 23 differentiation of osteoblasts and their regulation and protection from hyperglycaemia (Yan,
- 24 Li 2013), and metformin is considered to have beneficial effects on insulin resistance
- 25 (Wiernsperger, Bailey 1999).
- 26 We hypothesised that patients with type 2 diabetes treated with metformin may show a
- 27 reduced risk of OA compared to people with type 2 diabetes not so treated. To investigate
- 28 this, we conducted a longitudinal analysis using routinely recorded electronic health record
- 29 data.

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#### Methods

- 32 Study Design and Setting
- 33 This study used a cohort design using the Consultations in Primary Care Archive (CiPCA)
- database, an anonymised database of routinely recorded information from 13 general
- practices in North Staffordshire, UK (Porcheret et al. 2004, Jordan et al. 2007). Practices
- 36 undergo regular assessment, feedback and training on the quality of their morbidity
- 37 recording (Porcheret et al. 2004). Prevalence of consultation for musculoskeletal conditions
- has been shown to be similar to national and international databases (Jordan et al. 2013).
- 39 Practices contributing to CiPCA use the Read code system for recording morbidity as is most
- 40 common in UK primary care.
- 41 Study population
- To be eligible, patients had to be aged 40 or over and have had either a recorded diabetes
- diagnosis or diabetes treatment between January 2002 and December 2003 (the "baseline"
- 44 period"). Read codes for diabetes are available from our website www.keele.ac.uk/mrr.

Each patient's index date was defined as the first occurrence of a diagnosis of diabetes or prescription of a diabetic drug. Patients with a prior record of OA in the previous 2 years were excluded, as were patients with a record of type 1 diabetes (identified either through Read code or through linked consultation text). All eligible participants had a minimum of one year prior registration at their practice.

#### Exposure

Prescription information was available for all participants for their time in the study and at least 12 months prior to cohort entry. Those patients prescribed a drug for diabetes (BNF Chapter 6.1) would typically have multiple repeat prescriptions, and may switch between metformin and non-metformin treatments. A non-metformin prescription was defined as any other diabetic drug, or a diet and lifestyle advice (no drug) treatment only. Two approaches were conducted to investigate association of exposure to metformin with OA. The first analysis compared risk of future OA diagnosis based on baseline treatment (metformin prescription or not in the baseline period 2002-2003).

The second analysis incorporated change in pharmacological treatment of diabetes (metformin versus non metformin) over time. Patients not prescribed metformin in the baseline period but later prescribed metformin were deemed to be exposed to metformin from the date of the first such prescription. If a patient prescribed metformin was then not recorded as having a metformin prescription for 6 months at any subsequent point during follow-up, metformin exposure was deemed to have ended 28 days after the last recorded prescription If a prescription was recorded within 6 months of a prior prescription, metformin exposure was deemed to have continued uninterrupted. This is consistent with

- 67 previous work within our Research Institute and with guidelines that GPs should prescribe a
- 68 maximum of a 28 day supply of medication per prescription.
- 69 Outcome
- 70 The primary outcome of interest was the first occurence of an OA diagnosis during the
- 71 follow-up period, defined by Read code N05 "Osteoarthritis and allied disorders" and all
- 72 child codes. Follow-up continued to the end of 2011, the end of patient registration at their
- 73 practice, end of practice records in CiPCA, or the first record of OA.
- 74 Covariates
- 75 Covariates considered to be potential confounders of the relationship of metformin with OA
- diagnosis included age at index date, gender, GP practice, neighbourhood deprivation, and
- 77 comorbidity. Comorbidity was defined as number of different prescription drugs (based on
- 78 British National Formulary codes) prescribed during the baseline period and categorised into
- 79 4 groups; 0-5, 6-9, 10-13, and 14+, based on quartiles. This measure has been shown to be
- an efficient measure of comorbidity for healthcare use (Perkins et al. 2004).
- 81 Measurement of neighbourhood deprivation was based on the Index of Multiple
- 82 Deprivation (IMD) 2007, a small area-level measure of deprivation across England
- 83 (Department for Communities and Local governments 2007). This variable was categorised
- based on quintiles, the first category representing the most deprived in the population, and
- 85 the fifth category representing the least deprived.
- 86 Statistical Analyses
- 87 Cox proportional hazards regression models were fitted with Gamma frailty term. This is
- 88 essentially a random effects model to address variability in outcomes across patients (i.e.

different underlying frailty) related to unobserved covariates (Hougaard 1995). The shared frailty term in this case assumes that the frailty is common to patients within the same practice.

The proportional hazards assumptions were checked for both models fitted, and sensitivity analyses were conducted to test the robustness of the results to the distributional assumptions placed on the random effect. In place of a Gamma distribution, the commonly used Gaussian frailty term was added to the model (Yashin 2001).

All analyses were completed using R version 3.2.2 through R studio version 0.99.473 for Windows.

### Results

54,006 patients aged forty and over were registered at the 13 CiPCA practices in 2002. There were 4164 patients with a record of diabetes in 2002 or 2003. Of these 133 were excluded due to having a record of type 1 diabetes; 98 due to having no consultation information recorded during follow-up; 712 due to having a diagnosis of OA prior to their start date in the study; and a further 4 were removed due to having a diagnosis of OA on their index date. The remaining 3217 patients were eligible to be included in the analysis.

Baseline exposure analysis

Initially patients were split into treatment groups based on prescriptions received during the baseline period. There were 1838 (57.13%) patients prescribed metformin, and 1379 (42.87%) not prescribed metformin; 13.92% of those in the non-metformin group were on

- lifestyle and diet changes only, whilst the remaining 86.08% received a prescription for
- 111 another anti-diabetic drug.
- 112 Those prescribed metformin at baseline tended to be younger (mean age 64.08 [SD: 11.33]
- years versus 68.64 [SD: 11.90) but were similar in terms of gender, deprivation and median
- number of other prescription drugs during baseline (table 1).
- 115 Median follow-up was 8.50 (IQR: 4.08, 9.86) years for those prescribed metformin, and 7.63
- 116 (IQR: 2.98, 9.47) for those not prescribed metformin.
- 117 347 (18.88%) of those prescribed metformin had a diagnosis of OA during follow-up
- (incidence: 301.26; 95% CI: (271.17, 334.69) per 10,000 person years); 244 (17.69%) of those
- not prescribed metformin at baseline had a diagnosis of OA (314.55/10,000; 95% CI: 277.46,
- 120 356.61)).
- 121 There was no association of baseline prescription of metformin with OA (unadjusted HR:
- 122 0.97, (95% Cl: 0.87, 1.10), adjusted HR: 1.02 (95% Cl: 0.91, 1.15))(table 2).
- 123 Age (HR: 1.01 per year, (95% CI: 1.01, 1.02)), female gender (HR: 1.28, (95% CI: 1.09, 1.52)),
- and more prescription drugs 14+ versus 0-5 (HR: 2.18, (95% CI: 1.71, 2.79)) were associated
- with OA diagnosis during follow-up, whereas deprivation was not associated with OA
- diagnosis. The gamma frailty term was significant, indicating significant heterogeneity
- 127 between GP practices.
- One practice had only 36 registered patients in the analysis and appeared to violate the
- proportional hazards assumption. Its removal from the analysis did not change the findings.
- 130 Changing the gamma frailty term to a Gaussian did not substantially change the hazard
- 131 ratios.

132 Time-varying analysis133 2289 (71.11%) patients h

2289 (71.11%) patients had a metformin prescription at some point during follow-up; 196

(8.56%) of these patients received a diagnosis of OA whilst they were on a metformin

135 prescription.

2885 (89.62%) patients had a period of follow-up when they were not prescribed metformin, of which 395 (13.69%) were diagnosed with OA whilst they were on a non-metformin prescription.

Prescription of metformin (allowing exposure to vary over time) was not associated with a new OA diagnosis (unadjusted HR 0.93 (95% CI: 0.78, 1.10), adjusted HR 0.98 (95% CI: 0.82, 1.16). There was a similar relationship with OA for gender, age, deprivation score, and number of recorded prescriptions as in the analysis conducted on the baseline data (table 2).

The addition of the random effect was again significant.

The proportional hazards assumption was checked and satisfied for this model. Using a Gaussian rather than Gamma frailty term did not substantially change the estimated HRs.

## Discussion

In this cohort of diabetes patients with up to 10 years of follow-up, no significant association was found between prescription of metformin and diagnosis of OA.

Consistent with previous literature we identified an increase in risk with increasing age and an increased risk for females. We also identified a dose response relationship with number of other prescription drugs as a marker of comorbidity and risk of OA.

Diabetes has been shown to have an independent association with OA (5). This is the first population-based cohort study to examine whether metformin, a common treatment for diabetes, can have a protective effect against OA in those with diabetes. A major strength of this project is the large sample size from a primary care population database that has been found to give similar results for prevalence of musculoskeletal conditions compared to national databases (Jordan et al. 2013). The findings are therefore likely to be generalisable to the UK as a whole. There is variability between clinicians in diagnosing and recording OA, with some preferring a non-specific 'joint pain' term (Jordan et al. 2016). We have used only the OA diagnostic term in this study and so may have under-ascertained cases of OA but are likely to have included patients with more severe joint pain (Jordan et al. 2007). 27.2% patients had no recorded type associated with a diabetes diagnosis but were assumed to have type 2. This may be have led to a degree of misclassification but there is no reason to believe this should have biased the results. We did not assess OA in specific joints. Since the metabolic phenotype of OA has been suggested to predominantly affect the hand, knee, and generalised OA (Bijlsma, Berenbaum & Lafeber 2011), further work to examine the effect of metformin on site-specific OA would be appropriate. Patients may have had a

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Although the estimated association between metformin and OA was adjusted for potential confounders such as age and gender, we lacked information for this analysis about other pertinent covariates such as BMI. As the prescription of metformin in clinical practice is

previous diagnosis of OA more than 2 years before the index date, and so in some cases we

may have identified new consulting episodes of OA rather than first ever diagnosis.

linked to increased BMI, the lack of adjustment for BMI may hide any association of 175 176 metformin with reduced risk of OA. 177 Similarly any effect of metformin may plausibly be linked to dosage and duration which we 178 have not investigated here. In conclusion, this study has not identified evidence of an association of metformin with OA 179 but further research should assess the effects of dosage and duration on treatment, 180 181 incorporate BMI, and ascertain associations with site-specific OA. Ethics approval 182 Ethical approval for CiPCA was granted by the North Staffordshire Research Ethics 183 Committee. 184 <u>Acknowledgements</u> 185 186 The Keele GP Research Partnership, the Informatics team at the Arthritis Research UK 187 Primary Care Centre, and the National Institute for Health Research (NIHR). Financial Support 188 189 CiPCA database is funded by the North Staffordshire Primary Care Research Consortium and 190 Keele University Research Institute for Primary Care and Health Sciences. 191 Lauren Barnett is funded by a National Institute for Health Research (NIHR) Research Methods Fellowship. This project presents independent research funded by the National 192 Institute for Health Research (NIHR). The views expressed are those of the author(s) and 193 194 not necessarily those of the NHS, the NIHR or the Department of Health.

- 195 <u>Conflict of Interest</u>
- 196 The authors have declared no conflicts of interest.

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