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7 Article type : Original Article

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- 10 Title: Rheumatological conditions as risk factors for self-harm: A retrospective cohort study
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This article has been accepted for publication and undergone full peer review but has not been through the copyediting, typesetting, pagination and proofreading process, which may lead to differences between this version and the <u>Version of Record</u>. Please cite this article as <u>doi:</u> 10.1002/ACR.24345

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28	c.d.mallen@keele.ac.uk
29	
30	Word count: 2899
31	Keywords:
32	ankylosing spondylitis; fibromyalgia; osteoarthritis; rheumatoid arthritis; self-harm
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43	
44	Financial disclosure
45	This study was funded by the National Institute for Health Research (NIHR) School for
46	Primary Care Research (Funding Round 13, Grant Reference 349) and supported by the
47	National Institute for Health Research (NIHR) Applied Research Centre (ARC) West
48	Midlands. The views expressed are those of the author(s) and not necessarily those of the
49	NIHR or the Department of Health and Social Care.
50	
51	CDM is funded by the NIHR Applied Research Collaboration (ARC) West Midlands, the NIHR
52	School for Primary Care Research and a NIHR Research Professorship in General Practice
53	(NIHR-RP-2014-04-026). ZP is funded by the NIHR [Clinician Scientist Award (CS-2018-18-
54	ST2-010)/NIHR Academy]. CCG is part funded by NIHR ARC West Midlands. The study
55	sponsors had no role in study design; in the collection, analysis, and interpretation of data; in
56	the writing of the report; and in the decision to submit the paper for publication.
57	Abatra at

59 To examine the risk of self-harm in rheumatological conditions 60 Methods 61 Retrospective cohort study using data from the Clinical Practice Research Datalink. Patients 62 with ankylosing spondylitis, fibromyalgia, osteoarthritis or rheumatoid arthritis were 63 identified between 1990–2016 and matched to patients without these conditions. Incident 64 self-harm was defined by medical record codes following a rheumatological diagnosis. 65 Incidence rates (per 10,000 person-years(PY)) were reported for each condition, both 66 overall and year-on-year (2000-2016). Cox regression analysis determined risk (hazard 67 ratios(HR), 95% confidence interval(CI)) of self-harm for each rheumatological cohort 68 compared to the matched unexposed cohort. Initial crude analysis was subsequently 69 adjusted and stratified by age and gender. Due to non-proportionality over time, 70 osteoarthritis was also stratified by disease duration(<1 year, $\ge 1-<5$ years, $\ge 5-<10$ years, 71 ≥10 years). 72 **Results** 73 Incidence of self-harm was highest in patients with fibromyalgia (25.12 (95%CI 22.45,28.11) 74 per 10,000 PY) and lowest for osteoarthritis (6.48 (6.20, 6.76)). There was a crude 75 association with each rheumatological condition and self-harm, except for ankylosing 76 spondylitis. Though attenuated, these associations remained after adjustment for 77 fibromyalgia (HR 2.06(95%CI 1.60,2.65)), rheumatoid arthritis (1.59(1.20,2.11)) and 78 osteoarthritis (1-<5years: 1.12 (1.01,1.24); \geq 5-<10 years: 1.35 (1.18,1.54)). Age and gender 79 were weak effect modifiers for these associations. 80 Conclusions 81 Primary care patients with fibromyalgia, osteoarthritis or rheumatoid arthritis (but not 82 ankylosing spondylitis) are at increased risk of self-harm compared to people without these 83 rheumatological conditions. Clinicians need to be aware of the potential for self-harm in 84 patients with rheumatological conditions (particularly fibromyalgia), explore mood and risk 85 with them, and offer appropriate support and management. 86 87

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Objective

106 Significant findings

- Patients with fibromyalgia, osteoarthritis or rheumatoid arthritis are at increased risk of
 self-harm.
- Age and gender do not act as strong effect modifiers in the relationship between these
 rheumatological conditions and self-harm
- 111 In these patients, healthcare professionals need to be aware of this risk and offer appropriate management.

Introduction

Prevention of self-harm, a key risk factor for suicide, is an international public health priority (1). However, data from the UK between 2001 to 2013 showed a significant rising trend in primary care consultations for self-harm for men and women (2). A potential approach to preventing self-harm is the targeting of high-risk groups, with the World Health Organization (WHO) recommending that healthcare professionals assess the potential for self-harm in patient groups with symptoms of chronic pain and depression (3). Though these factors can apply to patients with rheumatological conditions, it remains unclear whether patients in these disease groups are at an increased risk of self-harm.

Our interest lies in examining and comparing the risk of self-harm in specific rheumatological conditions (ankylosing spondylitis, fibromyalgia, osteoarthritis and rheumatoid arthritis). These are some of the most prevalent rheumatological conditions with known relationships to chronic pain and depression. Poor mental health, especially a history of a depressive disorder, is a strong risk factor for self-harm (4) and comorbid depression is frequently experienced by patients with these common rheumatological conditions, to varying degrees (5-9). Although pain is on the causal pathway for depression, pain is itself an independent risk factor for self-harm (10) and is commonly experienced by those with rheumatological conditions (11).

151 Previous research into the role of a range of chronic health conditions on the risk of self-152 harm has shown conflicting findings (12-15). However, investigation into the role of 153 rheumatological conditions as risk factors for self-harm is currently very limited, despite these conditions being among the leading causes of disability worldwide (16). Webb et al 154 155 reported an initial unadjusted increased risk of self-harm in patients with osteoarthritis, but 156 this association was not retained after adjustment (12). Singhal et al did find an increased 157 relative risk (RR) of self-harm for patients with inflammatory polyarthropathies (RR 1.4 158 (1.3–1.4) after the first year since diagnosis, but this group contained a wide mix of 159 rheumatological conditions (ICD-10 code: (M05–M09, M12–M14)) limiting the clinical 160 usefulness of such information (14). Our aim was to examine and compare the risk of self-161 harm in several rheumatological conditions from a primary care population over time and 162 consider the role of age and gender on such risk. 163 164 **Patients and Methods** 165 Study design 166 We undertook a matched retrospective cohort study using Read-coded patient consultation 167 data (the clinical coding system used in UK primary care) from the Clinical Practice Research 168 Datalink (CPRD), a database of anonymised primary care records covering around 7% of the 169 UK population (17). This provides both coded consultation and prescription data and is 170 representative of the UK population, with regards to age, sex and ethnicity (17). 171 172 Study Population 173 We identified patients aged ≥18 years with one of four rheumatological conditions 174 (ankylosing spondylitis, fibromyalgia, osteoarthritis or rheumatoid arthritis) between 1st 175 January 1990 and 31st December 2016. Each exposed cohort was identified by specified 176 Read codes, identified and refined from an internal code list repository 177 (www.keele.ac.uk/mrr) and assigned an index date corresponding to the date of their 178 diagnosis. Where a patient had consulted for more than one of the rheumatological 179 conditions of interest, they were placed into the rheumatological cohort for which they had

first consulted. A single matched unexposed cohort was constructed to be a comparison

cohort for all rheumatological conditions. The unexposed cohort included individuals

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without a previous coded diagnosis of the rheumatological conditions of interest in their medical record, but no further exclusion was made based on the presence of other chronic disease. Cases were grouped by 10-year age range and gender and then within these categories' frequency matched to unexposed patients. Each control was assigned a pseudo-index date, generated at random from between their 18th birthday and the end of the study. All individuals were subsequently examined for a self-harm Read code, based on the Read code list used in previous CPRD research (2) (Code list available at www.keele.ac.uk/mrr upon request). Patients with a self-harm code prior to index date were excluded and incidence based only on the first self-harm code reported post-rheumatological diagnosis.

192 Statistical Analysis

Descriptive statistics were initially used to characterise the sample of each rheumatological condition, including age, gender, practice-level deprivation, BMI, smoking status, alcohol consumption and previous diagnosis for anxiety and/or depression (the latter four defined by the closest value recorded before the index date). Incidence rates of self-harm per 10,000 person-years (PY) were determined for each rheumatological condition from 01/01/1990 to 31/12/2016 and year-on-year incidence trends were reported for 01/01/2000 to 31/12/2016 (reduced time-period used due to low incidence in earlier years). Patients contributed data after the latest of three events: (1) the study start date, (2) the date they registered at a participating practice, plus 6 months or (3) the date at which the practice was adjudged to reach internal quality standards, known as the 'up-to-standard' date.

Using Cox proportional-hazards regression analysis over the full time-period (1990-2016), crude hazard ratios (HR) were initially reported with 95% confidence intervals (CIs) to examine the association between the presence of each of the four rheumatological conditions and the subsequent incidence of self-harm compared to the matched unexposed cohort. Adjusted analysis was then undertaken, accounting for age, gender, practice-level deprivation, BMI, smoking status, alcohol consumption, anxiety and depression. Cases with missing data for smoking, BMI and alcohol consumption were included within analysis using a missing category approach. Imputation was not considered sensible in this case, as data were unlikely to be missing at random (18, 19). Proportionality of hazards was examined for

213	each model using schoemerd's residuals. Where variables showed evidence of non-
214	proportionality, they were included as time varying covariates. Further analyses were
215	conducted for each rheumatological cohort, stratifying by median age of the relevant
216	exposed cohort and gender. We defined our dichotomized age subgroups as "younger" or
217	"older" patients. This study was approved by the CPRD Independent Scientific Advisory
218	Committee (reference number 18_018R3). Data were analyzed with Stata software (version
219	15.1, StataCorp, College Station, TX, USA). A two-sided p-value < 0.05 was considered for
220	statistical significance.
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222	Results
223	Sample characteristics
224	The number of cases in each rheumatological condition group identified were: ankylosing
225	spondylitis 10,484, fibromyalgia 17,546; osteoarthritis 410,384 and rheumatoid arthritis
226	23,205, with a matched unexposed cohort of the same size for each condition. Patient
227	cohorts with fibromyalgia, osteoarthritis or rheumatoid arthritis were predominantly female
228	(85.5%, 59.5% & 66.7% respectively) and the mean age across the four condition cohorts
229	(similar in their unexposed cohorts) ranged from the youngest of 47 years (fibromyalgia) to
230	the oldest at 65 years (osteoarthritis) (Table 1).
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232	Incidence of self-harm in rheumatological cohorts
233	Incidence of self-harm was highest in the fibromyalgia cohort at 25.12 (95%CI 22.45, 28.11)
234	per 10,000 PY, but lower at 11.37 (9.37, 13.81) for ankylosing spondylitis, 9.70 (8.32, 11.29)
235	for rheumatoid arthritis and 6.48 (6.20, 6.76) for osteoarthritis (Table 2). Over a 16-year
236	period (2000-2016), and despite fluctuations, the year-on-year incidence rates of self-harm
237	across each rheumatological condition changed little, with incidence rates in $2016\mathrm{similar}$ to
238	those in 2000 (Figure 1).
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240	Risk of self-harm in rheumatological conditions
241	The risk of self-harm in patients with osteoarthritis was not proportional over the study
242	period. As such, risk in this condition group was examined over four separate time-periods
243	of disease duration, ≤1 year, 1-<5 years, ≥5-<10 years and ≥10 years. There was a significant

244 crude association between fibromyalgia (HR 3.01 (95%CI 2.42, 3.76)), rheumatoid arthritis 245 (1.72 (1.34, 2.22)) and for the majority of categories of osteoarthritis patients (excluding 246 those with a disease duration of less than 1 year (1.07 (0.90, 1.26)) and subsequent self-247 harm compared to their matched unexposed counterparts. There was no association 248 between ankylosing spondylitis and self-harm compared to the unexposed matched cohort 249 (1.21 (0.90, 1.64)).250 251 After adjustment, the association in each cohort was attenuated. Patients with fibromyalgia 252 remained twice as likely to self-harm as those without a rheumatological condition and 253 rheumatoid arthritis patients had an increased self-harm risk of 59%. The increased risk of 254 self-harm in osteoarthritis with a disease duration of 1-<5 years and ≥5-<10 years was 12% 255 and 35% respectively; however, patients with osteoarthritis for ≥10 years no longer had a 256 significant risk of self-harm (1.17 (0.98, 1.40)) (**Table 2**). 257 258 Age-specific risk of self-harm in rheumatological conditions 259 After stratifying by median age, we found there to be a similar association between the risk 260 of self-harm across the younger and older strata for the rheumatoid arthritis cohort. 261 However, this association was only statistically significant for the younger cohort (Younger: 262 1.67 (1.18, 2.36), Older: (1.52 (0.92, 2.52)). In the fibromyalgia cohort, both age categories 263 experienced a significant increase in risk of self-harm with older patients experiencing a 264 slightly greater risk (Younger: 2.28 (1.66, 3.13), Older 2.58 (1.67, 3.99)). There was a 265 somewhat mixed picture for osteoarthritis, with some younger patients seeing an increase in 266 risk (1-<5 years: 1.18 (1.03, 1.34); \geq 5-<10 years: 1.34 (1.13, 1.58), as did older patients with 267 osteoarthritis (1.22 (1.07, 1.38), across all disease durations). There remained no increased 268 risk of self-harm in patients with ankylosing spondylitis when stratified by age (**Table 3**). 269 270 Gender-specific risk of self-harm in rheumatological conditions 271 There were similar levels of increased risk of self-harm across the genders in those with 272 osteoarthritis, rheumatoid arthritis or fibromyalgia compared to those without a 273 rheumatological condition. Men and women who had had osteoarthritis for between ≥5-<10 274 years had a significantly increased risk of self-harm of 42% and 35% respectively. However,

though the risk estimates for men with rheumatoid arthritis or fibromyalgia were similar to that of women with the same condition, the increase was not statistically significant in either gender (fibromyalgia men: 2.16 (0.91, 5.31); rheumatoid arthritis men 1.63 (0.98, 2.68)) (**Table 4**).

Discussion

We found that primary care patients with rheumatological conditions of fibromyalgia, rheumatoid arthritis or osteoarthritis are at increased risk of self-harm compared to matched unexposed patients, that this risk varies across the three conditions, but that age and gender are weak effect modifiers. In contrast there was no association between ankylosing spondylitis and self-harm. Finally, the year-on-year incidence of self-harm in these patient groups remained relatively consistent from 2000 to 2016.

Overall, patients with fibromyalgia had the greatest risk of self-harm with a 2-fold increase compared to the unexposed cohort. This propensity to self-harm was greater than that seen for the rheumatoid arthritis and osteoarthritis cohorts and may relate to different effects of similar mechanisms. Patients with fibromyalgia have been shown to experience more depression (6, 7, 20) and poorer health outcomes, particularly bodily pain (21) and fatigue (22), than patients with osteoarthritis and rheumatoid arthritis. Furthermore, in contrast to osteoarthritis and rheumatoid arthritis, psychological factors play a major role in the origin of this disorder as both physical and childhood trauma are associated with fibromyalgia and self-harm (4, 23). Though the risk of self-harm in different rheumatological conditions has not been previously described, a meta-analysis by Li et al did find the prevalence of suicidal ideation (of which self-harm is a risk factor (24)) to be greater in patients with fibromyalgia compared to the other rheumatological conditions (systemic lupus erythematosus and osteoarthritis) (25). This suggests the greatest risk we observed in patients with fibromyalgia is plausible.

Our initial findings on osteoarthritis were in-line with those of Webb et al, who examined the risk of self-harm in several physical illnesses using CPRD (12) and found an unadjusted increase in the risk of self-harm in patients with osteoarthritis. However, in contrast, we

found that after adjustment the risk of self-harm remained for certain disease duration categories. In our study, the risk remained increased in those having had osteoarthritis for between 1 and 10 years. Our contrasting findings to Webb et al are potentially due to our increased sample size from a much longer time-period of the CPRD, and thus greater statistical power. However, it is clear that the risk of self-harm in osteoarthritis is relatively modest, especially compared to fibromyalgia and rheumatoid arthritis. This may reflect a different experience in terms of pain duration, frequency and location; osteoarthritis is typically a more localised pain condition, unlike the generalized pain and systemic symptoms caused by fibromyalgia and rheumatoid arthritis, and therefore may have less impact on patients' lives leading to self-harm behaviour. Furthermore, the violation of the proportional hazards assumption in our study suggests that the risk of self-harm is influenced by disease duration, this may have been masked in the study of Webb et al (12) where data were from a case-control study and so the effect of disease duration could not be assessed. Though there is mixed evidence around the association between disease duration and progression of osteoarthritis (26), we found that those with newly diagnosed osteoarthritis (≤1 year) or those with long-standing disease (≥10 years) were not at significantly increased risk of self-harm. This could be due to patients with a recent diagnosis not yet having a prolonged negative experience of osteoarthritis, which could be a contributing factor to self-harm, or those with very long-standing osteoarthritis having developed coping mechanisms (27).

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Despite crude analysis demonstrating an initial association, adjusted analysis did not show ankylosing spondylitis to be a risk factor for self-harm. These findings may be related to this cohort having the greatest proportion of men, with a reduced incidence of self-harm compared to women (2). However, our matching, adjustment and stratification should have countered for the majority of confounding for gender. Though not examining self-harm, Wu et al. found no association between ankylosing spondylitis and subsequent suicide ideation or attempt (9). Although Wu et al., used a different group of patients, their sample was from a large population-based cohort which suggests our findings are conceivable and not just an artifact of our ankylosing spondylitis cohort being the smallest sample.

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Overall, we found age and gender to be weak effect modifiers of the relationship between all rheumatological cohorts and subsequent self-harm. For patients with fibromyalgia or rheumatoid arthritis, we found little difference in the strength of association between younger and older, or male and female patients and the subsequent risk of self-harm. However, this finding is tempered by relatively small participant numbers, which meant that the association was not statistically significant for men with fibromyalgia or rheumatoid arthritis and older patients with rheumatoid arthritis. Therefore, further study on the role of these factors is required to clarify this association.

Our findings suggest that primary care clinicians, rheumatologists and allied health professionals need to be especially aware of the potential for self-harm in people with fibromyalgia and rheumatoid arthritis, exploring mood and risk, and offering appropriate support and management. Interventions to reduce this serious comorbidity in rheumatological patients are important because, as our incident data show, little has changed in the proportion of patients who self-harm in recent years.

With regard to the strengths and limitations of our work, this is the first study to examine several rheumatological conditions as risk factors for incident self-harm. Our use of a large UK primary care dataset has allowed us to examine the incidence of self-harm over a 26-year time-period, year-on-year incidence trends over a 16-year time-period and examine the role of age and gender on risk of self-harm. Our analysis also takes account of clinically recorded depression, a key risk factor for self-harm. However, several limitations to our work need to be considered. Across these rheumatological conditions of interest, pain is likely to be an important contributory factor in the risk of self-harm (28), but could not be determined from consultation record data. Furthermore, it remains unclear to what extent use of medications influence the risk of self-harm, though the potential causal mechanism of such medications means this would be difficult to disentangle. We also found large proportions of missing data for BMI, smoking and alcohol consumption within the CPRD dataset, particularly in the unexposed patients. As such data is not "missing at random", we were unable to impute for these variables. However, we included missing data as a separate category in the Cox model and reported the extent of missing data in descriptive tables to ensure transparency. There

may also be some residual confounding. Firstly for psychological comorbidities, though our adjustment for anxiety and depression will have covered the majority of these, and secondly for additional rheumatological comorbidities, where patients went on to develop a further rheumatological condition of interest (most likely the common condition of OA) after they had been grouped into a cohort based on their first rheumatological consultation. However, the small increased risk of self-harm in patients with OA that we found would suggest that any confounding would have a minimal impact on the risk experienced by those with the other rheumatological conditions of interest. Our original intention within this analysis had also been to examine the risk of suicide, as well as self-harm, in these rheumatological conditions. However, available samples were too small to conduct such analysis. The available sample also proved problematic when stratifying by gender. As the majority of included participants were female, stratification by gender lead to small numbers of events in the male cohorts (e.g. only 35 incidents of self-harm in men with fibromyalgia). However, stratification remains justified, as it highlights the role of age and gender.

In conclusion, patients with rheumatological conditions have increased risk of self-harm compared to matched unexposed patients, but age and gender do not act as strong effect modifiers. The incidence of self-harm in these conditions has remained relatively consistent over the last decade and a half and therefore clinicians should be vigilant, explore mood, assess risk, and offer appropriate support and management, especially to patients with fibromyalgia.

Acknowledgments

Data

This study is based in part on data from the Clinical Practice Research Datalink obtained under licence from the UK Medicines and Healthcare products Regulatory Agency. The data is provided by patients and collected by the NHS as part of their care and support. The

interpretation and conclusions contained in this study are those of the author/s alone in ALL publications.

The copyright of the morbidity definitions lists (©2014) used in this publication is owned by Keele University, the development of which was supported by the Primary Care Research Consortium. The authors would like to acknowledge Keele University's Prognosis and Consultation Epidemiology Research Group who have given us permission to utilise the morbidity definitions (©2014). For access/details relating to the morbidity definitions lists (©2014) please go to www.keele.ac.uk/mrr

Competing interests: The authors have declared that no competing interests exist.

Author contributions

JAP was involved with conceptualisation of the project idea, the subsequent acquisition of funding, overall study design and the undertaking and interpretation of analysis. ZP came up with the initial idea, was a funding co-applicant and involved in overall study design. RW was a funding co-applicant, curated the data and conducted analysis. AAS was a funding co-applicant and supported RW in data curation. CCG was involved with conceptualization of the project idea and the subsequent acquisition of funding. SM & RB provided support to JAP in undertaking data analysis. TS and AS were co-applicants in the acquisition of funding and CDM was involved with conceptualization of the project idea, the funding application and overall study design. All authors were involved in manuscript writing, have read, and approved the final version.

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Table 1: Characteristics of patients with rheumatological conditions and their matched cohorts (1990-2016)

Factor Ankylosing		g Spondylitis	Fibromyalgia Osteo		Osteoa	arthritis	Rheumat	oid Arthritis
	Exposed (%)	Unexposed (%)	Exposed (%)	Unexposed (%)	Exposed (%)	Unexposed (%)	Exposed (%)	Unexposed (%)
N	10,484	10,484	17,546	17,546	410,384	410,384	23,205	23,205
Mean age at index (SD)	46.9 (15.6)	46.6 (15.9)	46.9 (12.1)	46.6 (12.7)	64.6 (12.7)	64.2 (13.2)	56.7 (15.2)	56.4 (15.6)
Males	4,376 (41.7)	4,376 (41.7)	2,552 (14.5)	2,552 (14.5)	166,241 (40.5)	166,241 (40.5)	7,728 (33.3)	7,728 (33.3)
Median years follow-up (IQR)	7.4 (3.8. 12.2)	5.3 (2.7, 10.6)	5.8 (3.0, 9.9)	5.7 (2.7, 11.3)	6.8 (3.6, 11.0)	5.9 (2.9, 11.0)	6.2 (3.1, 10.5)	5.9 (2.8, 11.2)
Deprivation status								
Q1 (Least deprived)	1,986 (18.9)	1,917 (18.3)	3,002 (17.1)	3,218 (18.3)	72,018 (17.6)	73,347 (17.9)	3,774 (16.3)	4,186 (18.1)
Q2	1,806 (17.2)	1,782 (17.0)	2,960 (16.9)	3,227 (18.4)	71,941 (17.5)	72,068 (17.5)	4,070 (17.5)	4,042 (17.4)
Q3	2,109 (20.1)	2,190 (20.9)	3,337 (19.0)	3,571 (20.4)	79,179 (19.3)	85,833 (20.8)	4,752 (20.5)	4,824 (20.8)
Q4	1,925 (18.4)	2,259 (21.5)	3,786 (21.6)	3,766 (21.5)	85,826 (20.9)	90,589 (22.1)	4,892 (21.1)	5,041 (21.7)
Q5 (Most deprived)	2,658 (25.4)	2,336 (22.3)	4,461 (25.4)	3,764 (21.4)	101,420 (24.7)	89,047 (21.7)	5,717 (24.6)	5,112 (22.0)
BMI (kg/m ²)								
Underweight (<18.5)	164 (1.5)	123 (1.2)	286 (1.6)	218 (1.2)	3,013 (0.7)	3,587 (0.9)	389 (1.7)	234 (1.0)
Healthy weight (<18.5-24.9)	3,707 (35.4)	2,294 (21.9)	5,602 (31.9)	4,371 (24.9)	99,869 (24.3)	79,731 (19.4)	7,491 (32.3)	4,920 (21.2)
Overweight (25.0-29.9)	2,800 (26.7)	1,587 (15.1)	4,983 (28.4)	2,638 (15.0)	136,070 (33.2)	72,724 (17.7)	6,600 (28.4)	3,960 (17.1)
Obese (>30.0)	1,595 (15.2)	886 (8.4)	4,817 (27.5)	1,747 (10.0)	104,776 (25.5)	37,063 (9.2)	4,474 (19.3)	2,137 (9.2)
Missing	2,218 (21.2)	5,594 (53.4)	1,858 (10.6)	8,572 (48.9)	66,656 (16.3)	216,739 (52.8)	4,251 (18.3)	11,954 (51.5)
Smoking								
Never/Ex smoked	6,680 (63.7)	4,212 (40.2)	12,163 (69.3)	7,696 (43.9)	308,795 (75.2)	170,633 (41.6)	15,332 (66.1)	9,660 (41.6)
Current smoker	2,540 (24.2)	1,494 (14.2)	4,636 (26.4)	2,416 (13.8)	63,869 (15.6)	48,963 (11.9)	5,566 (24.0)	3,0.37 (13.1)
Missing	1,264 (12.1)	4,778 (45.6)	747 (4.3)	7,434 (42.4)	37,720 (9.2)	190, 788 (46.5)	2,307 (9.9)	10,508 (45.3)
Alcohol consumption								
Never/Ex-drinker	1,493 (14.2)	942 (9.0)	4,104 (23.4)	1,735 (9.9)	72,098 (17.6)	42,065 (10.2)	4,229 (18.2)	2,225 (9.6)
Current 1-9	4,923 (47.0)	3,000 (28.6)	9,355 (53.3)	5,824 (33.2)	200,504 (48.9)	116,325 (28.4)	11,178 (48.2)	6,891 (29.7)
Current >=10	1,6223 (15.5)	869 (8.5)	1,692 (9.6)	1,236 (7.0)	66,935 (16.3)	32,979 (8.0)	3,120 (13.5)	1,891 (8.1)
Missing	2,445 (23.3)	5,646 (53.9)	2,395 (13.7)	8,751 (49.9)	70,847 (17.2)	219,016 (53.4)	4,678 (20.1)	12,198 (52.6)
Anxiety	1,757 (16.8)	879 (8.4)	5,944 (33.9)	1,865 (10.6)	67,894 (16.5)	36,568 (8.9)	3,385 (14.6)	2,229 (9.6)
Depression	2,356 (22.5)	1,263 (12.1)	8,784 (50.1)	2,786 (15.9)	90,354 (22.0)	51,631 (12.6)	5,035 (21.7)	3,064 (13.2)

Table 2: Risk of self-harm associated with each rheumatological condition

		Exposed	Ţ	Jnexposed	Hazard ratios (95% CI)		
Condition	n Incidence rate, per		n	Incidence rate, per	Crude	Adjusted*	
		10,000 (95% CI)		10,000 (95% CI)			
Total							
Ankylosing Spondylitis	102	11.37 (9.37, 13.81)	72	9.57 (7.60, 12.06)	1.21 (0.90, 1.64)	0.95 (0.69, 1.31)	
Fibromyalgia	303	25.12 (22.45, 28.11)	108	8.24 (6.82, 9.95)	3.01 (2.42, 3.76)	2.06 (1.60, 2.65)	
Osteoarthritis	2,060	6.48 (6.20, 6.76)	1,528	5.02 (4.77, 5.28)			
≤1 year	277	6.75 (6.00,7.60)	260	6.34 (5.61, 7.16)	1.07 (0.90, 1.26)	1.02 (0.86, 1.21)	
1-<5 years	844	6.36 (5.95, 6.80)	664	5.36 (4.97, 5.78)	1.19 (1.07, 1.31)	1.12 (1.01, 1.24)	
≥5-<10 years	629	6.72 (6.22, 7.27)	376	4.46 (4.03, 7.27)	1.51 (1.33, 1.72)	1.35 (1.18, 1.54)	
10≥ years	310	6.11 (5.47, 6.83)	228	4.13 (3.63, 4.70)	1.47 (1.24, 1.74)	1.17 (0.98, 1.40)	
Rheumatoid Arthritis	165	9.70 (8.32, 11.29)	96	5.52 (4.52, 6.75)	1.72 (1.34, 2.22)	1.59 (1.20, 2.11)	

^{*}adjusted for age, BMI, smoking status, alcohol consumption, anxiety, depression and practice-level deprivation. **Bold** = statistically significant (p<=0.05).

Table 3: Age-specific risk of self-harm associated with each rheumatological condition

		Exposed		N	Non-exposed	Hazard ratios (95% CI)	
Condition	Total incidents of self-harm	n	Incidence rate, per 10,000 (95% CI)	n	Incidence rate, per 10,000 (95% CI)	Crude	Adjusted*
Ankylosing Spondylitis							
<45.3 years	131	78	17.63 (14.12, 22.01)	53	15.22 (11.62,19.92)	1.21 (0.85, 1.71)	0.97 (0.68, 1.40)
>= 45.3 years	43	24	5.28 (3.54, 7.88)	19	4.71 (3.00, 7.38)	1.13 (0.62, 2.06)	0.88 (0.46, 1.68)
Fibromyalgia							
<46.7 years	282	205	34.72 (30.27, 39.81)	77	12.30 (9.83, 15.37)	2.77 (2.13, 3.61)	2.28 (1.66, 3.13)
>=46.7 years	129	98	15.92 (13.06, 19.40)	31	4.52 (3.18, 6.43)	3.56 (2.38, 5.32)	2.58 (1.67, 3.99)
Osteoarthritis							
<64.6 years							
≤1 year	380	192	9.36 (8.13, 10.78)	188	8.99 (7.79, 10.37)	1.04 (0.85, 1.27)	1.01 (0.83, 1.24)
1-<5 years	985	561	8.25 (7.59, 8.96)	424	6.53 (5.94, 7.19)	1.26 (1.11, 1.43)	1.18 (1.03, 1.34)
≥5-<10 years	618	393	7.58 (6.86, 8.37)	225	4.73 (6.86, 8.37)	1.60 (1.36, 1.88)	1.34 (1.13, 1.58)
10≥ years	340	194	5.86 (5.09, 6.75)	146	4.06 (3.46, 4.78)	1.43 (1.15, 1.77)	0.99 (0.80, 1.25)
>=64.6 years							
≤1 year	157	85	4.14 (3.35, 5.13)	72	3.58 (2.84, 4.51)	1.16 (0.84, 1.58)	
1-<5 years	523	283	4.38 (3.89, 4.92)	240	4.07 (3.58, 4.61)	1.07 (0.90, 1.27)	1 22 (1 07 1 20)†
≥5-<10 years	387	236	5.66 (4.98, 6.42)	151	5.66 (4.98, 6.43)	1.38 (1.13, 1.69)	1.22 (1.07, 1.38) [†]
10≥ years	198	116	6.59 (5.49, 7.91)	82	4.26 (3.43, 5.29)	1.55 (1.17, 2.06)	
Rheumatoid Arthritis			-		-		
<57.0 years	184	120	13.20 (11.04, 15.79)	64	7.15 (5.60, 9.13)	1.83 (1.35, 2.48)	1.67 (1.18, 2.36)
>=57.0 years	77	45	5.68 (4.24, 7.60)	32	3.79 (2.68, 5.37)	1.50 (0.95, 2.37)	1.52 (0.92, 2.52)

^{*}adjusted for gender, BMI, smoking status, alcohol consumption, anxiety, depression and practice-level deprivation. **Bold** = statistically significant (p<=0.05). †Data met assumptions for proportionality and therefore stratification by time was not required

Table 4: Gender-specific risk of self-harm associated with each rheumatological condition

		Exposed		N	lon-exposed	Hazard ratios (95% CI)	
Condition	Total incidents of self-harm	n	Incidence rate, per 10,000 (95% CI)	n	Incidence rate, per 10,000 (95% CI)	Crude	Adjusted*
Ankylosing Spondylitis							
Men	71	42	11.59 (8.56, 15.68)	29	9.35 (6.50, 13.45)	1.25 (0.78, 2.01)	1.13 (0.69, 1.84)
Women	103	60	11.23 (8.72, 14.46)	43	9.73 (7.22, 13.12)	1.19 (0.80, 1.77)	0.83 (0.55, 1.27)
Fibromyalgia							
Men	48	35	18.34 (13.17, 25.54)	13	6.82 (3.96, 11.74)	2.98 (1.54, 5.75)	2.16 (0.91, 5.13)
Women	363	268	26.39 (23.42, 29.75)	95	8.48 (6.93, 10.37)	3.03 (2.40, 3.83)	2.44 (1.87, 3.18)
Osteoarthritis							
Men							
≤1 year	209	108	6.50 (5.38, 7.85)	101	6.08 (5.00, 7.39)	1.07 (0.82, 1.40)	1.06 (0.81, 1.40)
1-<5 years	568	320	6.01 (5.38, 6.70)	248	4.94 (4.36, 5.59)	1.22 (1.03, 1.44)	1.18 (0.99, 1.40)
≥5-<10 years	382	240	6.53 (5.76, 7.41)	142	4.17 (3.53, 4.91)	1.57 (1.27, 1.93)	1.42 (1.14, 1.76)
10≥ years	191	114	8.80 (4.83, 6.97)	77	3.50 (2.80, 4.38)	1.66 (1.24, 2.22)	1.34 (0.99, 1.81)
Women							
≤1 year	328	169	6.92 (5.96, 8.05)	159	6.51 (5.78, 7.61)	1.06 (0.86, 1.32)	1.08 (0.87, 1.35)
1-<5 years	940	524	6.60 (6.06, 7.19)	416	5.64 (5.13, 6.21)	1.17 (1.03, 1.33)	1.16 (1.02, 1.33)
≥5-<10 years	623	389	6.85 (6.20, 7.56)	234	4.65 (4.09, 5.29)	1.47 (1.25, 1.73)	1.35 (1.14, 1.59)
10≥ years	347	196	6.31 (5.49, 7.26)	151	4.55 (3.88, 5.33)	1.37 (1.11, 1.70)	1.07 (0.85, 1.34)
Rheumatoid Arthritis							
Men	82	53	9.81 (7.50, 12.84)	29	5.03 (3.49, 7.23)	1.89 (1.20, 2.99)	1.63 (0.98, 2.68)
Women	179	112	9.64 (8.01, 11.60)	67	5.77 (4.54, 7.33)	1.65 (1.22, 2.23)	1.56 (1.11, 2.20)

^{*}adjusted for age, BMI, smoking status, alcohol consumption, anxiety, depression and practice-level deprivation. **Bold** = statistically significant (p<=0.05). AS = Ankylosing Spondylitis, OA = Osteoarthritis, RA = Rheumatoid Arthritis

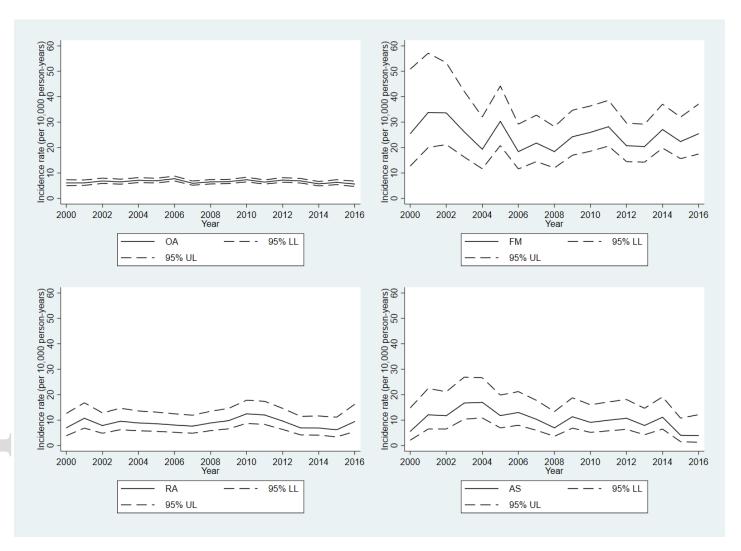


Figure 1 legend: Incidence rates per 10,000 person-years for osteoarthritis (OA), fibromyalgia (FM), rheumatoid arthritis (RA) and ankylosing spondylitis (AS). Solid line = rate, dashed line = 95% confidence intervals, lower and upper bands

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